# Chapter 8 <br> Scaling the Data from the TIMSS Advanced 2008 Assessments 

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### 8.1 Overview

The TIMSS Advanced 2008 goals of broad coverage of the advanced mathematics and physics curricula and of measuring trends across assessments necessitated the adoption of a complex matrix-sampling booklet design, ${ }^{1}$ with individual students responding to a subset of the advanced mathematics and physics items in the assessment, and not the entire assessment item pool. Given the complexities of the data collection and the need to have student scores on the entire assessment for analysis and reporting purposes, TIMSS Advanced relied on Item Response Theory (IRT) scaling to describe student achievement on the assessment and to provide accurate measures of trends from the previous assessments. The TIMSS IRT scaling approach used multiple imputation-or "plausible values"-methodology to obtain proficiency scores in advanced mathematics and physics for all students, even though each student responded to only a part of the assessment item pool. To enhance the reliability of the student scores, the TIMSS
advanced scaling combined student responses to the items they were administered with information about students' backgrounds, a process known as "conditioning."

This chapter describes the steps that produced scaled scores of student achievement in advanced mathematics and physics. First, it explains the process of reviewing item statistics to validate the statistical properties of the achievement items used in the TIMSS Advanced 2008 assessments. It then provides a general explanation of the methodology used to scale the TIMSS Advanced 2008 data, and describes how this approach was applied to the 2008 assessment data and to the data from the previous TIMSS Advanced 1995 study in order to measure trends in achievement. The TIMSS Advanced scaling was conducted by the TIMSS \& PIRLS International Study Center at Boston College, using software from Educational Testing Service. ${ }^{2}$ This chapter also provides a description of the scale anchoring methodology used to describe student performance at various points on the TIMSS Advanced mathematics and physics achievement scales, and the methodology used to estimate standard errors of the estimates published for TIMSS Advanced 2008.

### 8.2 Item Review

For TIMSS Advanced 2008, as in the TIMSS assessments at the fourth and eighth grades, the TIMSS \& PIRLS International Study Center conducted a review of a range of diagnostic statistics to examine and evaluate the psychometric characteristics of each achievement item in the ten countries that participated in TIMSS Advanced 2008. This review of item statistics was conducted before applying item response theory (IRT) scaling to the TIMSS Advanced 2008 achievement data to derive student achievement scores in advanced mathematics and physics for analysis and reporting. The review of item statistics played
a crucial role in the quality assurance of the TIMSS Advanced 2008 data, making it possible to detect unusual item properties that could signal a problem or error for a particular country. For example, an item that was uncharacteristically easy or difficult, or had an unusually low discriminating power, could indicate a potential problem with either translation or printing. Similarly, a constructed-response item with unusually low scoring reliability could indicate a problem with a scoring guide in a particular country. In the rare instances where such items were found, the country's translation verification documents and printed booklets were examined for flaws or inaccuracies and, if necessary, the item was removed from the international database for that country.

### 8.2.1 Statistics for Item Analysis

To begin the review process, the TIMSS \& PIRLS International Study Center computed item statistics for all 143 advanced mathematics and physics achievement items that were administered in the TIMSS Advanced 2008 assessments. The properties of the items in each of the ten countries that participated were then carefully reviewed. Exhibits 8.1 and 8.2 show actual samples of the statistics calculated for a multiple-choice and a constructed-response item, respectively.

For all items, regardless of format, statistics included the number of students that responded in each country, the difficulty level (the percentage of students that answered the item correctly), and the discrimination index (the point-biserial correlation between success on the item and a total score). ${ }^{3}$ Also provided was an estimate of the difficulty of the item using a Rasch one-parameter IRT model. The international means of the item difficulties and item discriminations served as guides to the overall statistical properties of the items. Statistics for each item were displayed alphabetically by country,

Exhibit 8.1 International Item Statistics for a Multiple-choice Item
Trends in International Mathematics and Science Study - TIMSS 2008 Assessment Results
International Item Statistics (Unweighted) - Review Version - 12 th Grade International Item Statistics (Unweighted) - Review Version - 12th Grade
For Internal Review Only: DO NOT CITE OR CIRCULATE
Mathematics: Calculus / Reasoning (MA23206-M4_06)
Label: Sign of derivative function
Type: MC Key: D

|  |  |  |  | Percentages |  |  |  |  |  |  | Point Biserials |  |  |  |  |  | RDIFF | Flags |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Country | N | Diff | Disc | Pct_A | Pct_B | Pct_C | Pct_D | Pct_E | Pct_OM | PCt_NR | PB_A | PB_B | PB_C | PB_D | PB_E | PB_OM |  |  |
| Armenia | 439 | 41.5 | 0.45 | 32.8 | 9.1 | 8.4 | 41.5 |  | 8.2 | 0.2 | -0.14 | -0.22 | -0.20 | 0.45 |  | -0.13 | -0.77 | F |
| Iran, Islamic Rep. of | 1195 | 59.7 | 0.54 | 13.6 | 5.4 | 9.5 | 59.7 | . | 11.7 | 0.1 | -0.28 | -0.10 | -0.22 | 0.54 | . | -0.26 | -0.78 | $\mathrm{F}_{-}^{-}$ |
| Italy | 1080 | 44.3 | 0.47 | 44.4 | 2.2 | 4.1 | 44.3 | . | 5.1 | 0.1 | -0.39 | -0.09 | -0.06 | 0.47 | . | -0.06 | -0.70 | $\mathrm{F}_{-}^{-}$ |
| Lebanon | 801 | 80.9 | 0.37 | 15.7 | 0.6 | 2.1 | 80.9 |  | 0.6 | 0.0 | -0.32 | -0.08 | -0.10 | 0.36 | . | -0.07 | -1.36 | E_F- |
| Netherlands | 763 | 87.5 | 0.25 | 5.4 | 1.3 | 5.6 | 87.5 | . | 0.1 | 0.3 | -0.17 | -0.05 | -0.16 | 0.25 | . | -0.03 | -1.94 | E_F_ |
| Norway | 966 | 52.3 | 0.38 | 28.8 | 8.6 | 8.1 | 52.3 | . | 2.3 | 0.0 | -0.33 | -0.03 | -0.09 | 0.38 | . | -0.08 | -1.15 | ${ }^{-} \mathrm{F}_{-}^{-}$ |
| Philippines | 2057 | 17.6 | 0.23 | 61.1 | 9.7 | 10.3 | 17.6 | . | 1.3 | 0.0 | -0.16 | -0.07 | 0.05 | 0.23 | . | -0.02 | 0.15 | $\mathrm{CH}_{-} \mathrm{F}_{-}^{-}$ |
| Russian Federation | 1588 | 68.1 | 0.54 | 24.5 | 0.9 | 5.5 | 68.1 | . | 1.0 | 0.0 | -0.42 | -0.08 | -0.22 | 0.54 | . | -0.11 | -0.64 | $\mathrm{H}_{-}^{-} \mathrm{F}_{-}^{-}$ |
| Slovenia | 1090 | 38.3 | 0.46 | 41.2 | 4.4 | 12.3 | 38.3 | . | 3.9 | 0.0 | -0.36 | -0.04 | -0.10 | 0.46 | . | -0.04 | -0.26 | H_F- |
| Sweden | 1141 | 67.3 | 0.48 | 17.0 | 4.8 | 9.6 | 67.3 | . | 1.3 | 0.3 | -0.35 | -0.11 | -0.19 | 0.48 | . | -0.11 | -2.01 | $\mathrm{E}_{-}^{-} \mathrm{F}_{-}$ |
| International Avg. | 1112 | 55.8 | 0.42 | 28.4 | 4.7 | 7.5 | 55.8 |  | 3.6 | 0.1 | -0.29 | -0.09 | -0.13 | 0.42 | . | -0.09 | -0.95 | F ${ }_{-}$ |

[^0]TIMSS \& PIRLS
International Study Center

Exhibit 8.2 International Item Statistics for a Constructed-response Item
Trends in International Mathematics and Science Study - TIMSS 2008 Assessment Results International Item Statistics (Unweighted) - Review Version - 12th Grade
For Internal Review Only: DO NOT CITE OR CIRCULATE Mathematics: Calculus / Reasoning (MA23157 - M5_06) Label: Find point of maxima and inflection
Type: CR Key: X

| Country | N | Diff | Disc | Pct_0 | Percentages |  |  |  | Point Biserials |  |  |  | RDIFF | Reliability |  |  | Flags |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | Pct_1 | Pct_2 | Pct_OM | t_NR | PB_0 | PB_1 | PB_2 | PB_OM |  | Cases | Score | Code |  |
| Armenia | 407 | 11.5 | 0.57 | 24.1 | 13.3 | 4.9 | 57.7 | 4.9 | -0.14 | 0.43 | 0.36 | -0.32 | 1.02 | 91 | 100.0 | 100.0 | $\mathrm{F}_{-}$ |
| Iran, Islamic Rep. of | 1132 | 33.4 | 0.68 | 27.0 | 23.8 | 21.6 | 27.7 | 6.8 | -0.25 | 0.13 | 0.61 | -0.41 | 0.52 | 294 | 100.0 | 98.3 |  |
| Italy | 1032 | 16.2 | 0.55 | 27.8 | 17.6 | 7.4 | 47.2 | 3.4 | -0.12 | 0.20 | 0.49 | -0.29 | 0.83 | 139 | 100.0 | 100.0 | $\mathrm{F}_{-}$ |
| Lebanon | 784 | 39.5 | 0.47 | 26.4 | 50.5 | 14.3 | 8.8 | 2.1 | -0.32 | 0.12 | 0.38 | -0.18 | 0.82 | 188 | 99.5 | 99.5 |  |
| Netherlands | 757 | 43.8 | 0.34 | 22.3 | 58.0 | 14.8 | 4.9 | 1.2 | -0.19 | 0.04 | 0.28 | -0.18 | 0.60 | 198 | 94.9 | 93.9 |  |
| Norway | 936 | 28.6 | 0.39 | 28.5 | 45.0 | 6.1 | 20.4 | 2.6 | -0.21 | 0.10 | 0.39 | -0.17 | 0.41 | 235 | 100.0 | 99.1 | E_F |
| Philippines | 2033 | 16.7 | 0.38 | 50.2 | 31.0 | 1.2 | 17.6 | 0.2 | -0.21 | 0.24 | 0.36 | -0.12 | 0.89 | 507 | 95.1 | 93.9 | $\mathrm{H}_{-} \mathrm{F}_{-}$ |
| Russian Federation | 1582 | 57.2 | 0.47 | 12.2 | 41.2 | 36.6 | 10.1 | 0.9 | -0.19 | -0.10 | 0.42 | -0.30 | 0.02 | 402 | 98.8 | 98.8 | E |
| Slovenia | 1058 | 20.1 | 0.43 | 44.6 | 31.8 | 4.3 | 19.4 | 1.9 | -0.19 | 0.26 | 0.32 | -0.21 | 1.04 | 283 | 100.0 | 99.3 | H_F |
| Sweden | 1112 | 20.7 | 0.54 | 46.8 | 21.8 | 9.8 | 21.7 | 4.1 | -0.31 | 0.21 | 0.46 | -0.24 | 0.36 | 289 | 96.9 | 94.5 | E_F- |
| International Avg. | 1083 | 28.8 | 0.48 | 31.0 | 33.4 | 12.1 | 23.5 | 2.8 | -0.21 | 0.16 | 0.41 | -0.24 | 0.65 |  | 98.5 | 97.7 |  |

[^1]together with the international average for each statistic in the last row of the exhibits.

Statistics displayed for multiple-choice items included the percentage of students that chose each response option, as well as the percentage of students that omitted or did not reach the item, and the point-biserial correlations for each response option. Statistics displayed for constructed-response items (which could have 1 or 2 score points) included the difficulty and discrimination of each score level. Constructed-response item displays also provided information about the reliability with which each item was scored in each country, showing the total number of double-scored cases, the percentage of code agreement between the scorers, and the percentage of score agreement between scorers.

The definitions and detailed descriptions of the statistics that were calculated are given below for the examples shown in Exhibits 8.1 and 8.2. The statistics were calculated separately for advanced mathematics and physics. The statistics are listed in order of their appearance in the item analysis outputs:

N : The number of students to whom the item was administered. If a student did not reach an item in the achievement booklet, the item was considered not administered for item analysis. ${ }^{4}$

Diff: The item difficulty is the average percent correct. For a 1 -point item, including multiple-choice items, it is the percentage of students providing a fully correct response to the item. For 2-point items, it is the average percentage of points; i.e., if all students scored 1 point on a 2-point item, then the average percent correct for such an item would be 50 percent. For this statistic, not-reached items were treated as not administered.

Disc: The item discrimination was computed as the correlation between a correct response to the item and the overall score on all of the advanced mathematics or physics items administered to a student. ${ }^{5}$ Items exhibiting good measurement properties should have a moderately positive correlation, indicating that the more able students get the item right, the less able get it wrong.

PCT_A, PCT_B, PCT_C, PCT_D, and PCT_E: Used for multiplechoice items only (see Exhibit 8.1). Each column indicates the percentage of students choosing the particular response option for the item (A, B, C, D, or E). Not-reached items were excluded from the denominator.

PCT_o, PCT_1, and PCT_2: Used for constructed-response items only (see Exhibit 8.2). Each column indicates the percentage of students scoring at the particular score level, up to and including the maximum score level for the item. Not-reached items were excluded from the denominator.

PCT_OM: Percentage of students who, having reached the item, did not provide a response. No reached items were excluded from the denominator.

PCT_NR: Percentage of students who did not reach the item. An item was coded as not reached when there was no evidence of a response to any subsequent items in the booklet and the response to the item preceding it was also omitted.
PB_A, PB_B, PB_C, PB_D, and PB_E: Used for multiple-choice items only. These columns show the point-biserial correlations between choosing each of the response options (A, B, C, D, or E) and the overall score on all of the advanced mathematics or physics items administered to a student. Items with good
psychometric properties have moderately positive correlations for the correct option and negative correlations for the distractors (the incorrect options).

PB_o, PB_1, and PB_2: Used for constructed-response items only. These columns present the point-biserial correlations between the score levels on the item ( 0,1, or 2 ) and the overall score on all of the mathematics or science items the student was administered. For items with good measurement properties, the correlation coefficients should increase from negative to positive as the score on the item increases.

PB_OM: The point-biserial correlation between a binary variable, indicating an omitted response to the item, and the overall score on all of the mathematics or physics items administered to a student. This correlation should be negative or near zero.

RDIFF: An estimate of the difficulty of an item based on a Rasch one-parameter IRT model applied the achievement data for a given country. The difficulty estimate is expressed in the logit metric (with a positive logit indicating a difficult item) and was scaled so that the average Rasch item difficulty across all items within each country was zero.

Reliability (Cases): To provide a measure of the reliability of the scoring of constructed-response items, items in approximately 25 percent of the test booklets in each country were independently scored by two scorers. This column indicates the number of responses that were double-scored for a given item in a country.

Reliability (Score): This column contains the percentage of agreement on the score value of the two-digit codes assigned by the two independent scorers.

Reliability (Code): This column contains the percentage of exact agreement on the two-digit codes assigned by the two independent scorers.

As an aid to the reviewers, the item-analysis displays included a series of flags signaling the presence of one or more conditions that might indicate a problem with an item. The following conditions were flagged:

- Item discrimination was less than 0.15 (flag D).
- Item difficulty was less than 25 percent for four-option multiplechoice items, or less than 20 percent for five-option multiple choice items (flag C).
- Item difficulty exceeded 95 percent (flag V).
- The Rasch difficulty estimate for a given country made the item either easier (flag E) or harder (flag H) relative to the international average Rasch difficulty of that item.
- The point-biserial correlation for at least one distractor in a multiple choice item was positive, or the estimated mean abilities across the score levels of a constructed-response item were not ordered (flag A).
- The percentage of students selecting one of the response options for a multiple-choice item or of one of the score values for a constructed-response item was less than 10 percent (flag F).
- Scoring reliability for agreement on the score value of a constructed-response item was less than 80 percent (flag R).

Although not all of these conditions necessarily indicated a problem, the flags were a useful way to draw attention to potential sources of concern.

In order to measure trends, TIMSS Advanced 2008 included items from the 1995 assessments. ${ }^{6}$ For these trend items, the review included an examination of changes in item statistics between the 1995 and 2008 administrations. An example is shown in Exhibit 8.3. The information in this exhibit is different from that presented in Exhibits 8.1 and 8.2, and includes countries' statistics from both the 1995 and 2008 assessments. In reviewing these item statistics, the aim was to detect any unusual changes in item properties between assessments that might indicate a problem in using the item to measure trends.

## Exhibit 8.3 International Item Statistics for a Trend Item

Trends in International Mathematics and Science Study - TIMSS Advanced 2008 Assessment Results Percent of Responses by Item Category (Mathematics) - Trend Items - Final Year of Secondary School

Mathematics: Geometry / Applying (MA13026A - M3_06A)
Label: Triangle abc/reflection
Type: CR Key: X

|  |  | N | 10 | 70 | 71 | 79 | 99 | $\begin{gathered} \text { INVA } \\ \text { LID } \end{gathered}$ | OMIT | NOT |  |  | \% | $\begin{aligned} & \text { 2.BOY } \\ & \text { Right } \end{aligned}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Year |  |  |  |  |  |  |  |  | $\begin{array}{r} \text { REACH } \end{array}$ | V1 | 1.GIRL \% Right |  |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| Italy | 1995 | 126 | 65.8 | 4.3 | 5.1 | 11.8 | 13.1 | 1.5 | 5.0 | 6.6 | 65.8 | 65.8 |  | 65.7 |
|  | 2008 | 1070 | 54.9 | 2.5 | 6.7 | 21.2 | 14.8 | 0.0 | 8.8 | 6.0 | 54.9 | 56.5 |  | 54.0 |
| Russian Federation | 1995 | 468 | 75.4 | 3.8 | 5.3 | 9.2 | 6.3 | 0.0 | 1.2 | 5.1 | 75.4 | 74.4 |  | 76.3 |
|  | 2008 | 1588 | 75.0 | 2.1 | 6.2 | 10.7 | 5.9 | 0.0 | 4.4 | 1.6 | 75.0 | 74.5 |  | 75.5 |
| Slovenia | 1995 | 452 | 69.2 | 1.4 | 0.3 | 21.9 | 7.3 | 2.0 | 2.2 | 3.1 | 69.2 | 66.7 |  | 71.6 |
|  | 2008 | 1083 | 69.7 | 7.6 | 2.7 | 16.9 | 3.1 | 0.0 | 0.9 | 2.2 | 69.7 | 68.7 |  | 71.2 |
| Sweden | 1995 | 244 | 39.5 | 10.0 | 1.6 | 25.4 | 23.5 | 0.4 | 13.6 | 9.5 | 39.5 | 25.4 |  | 47.1 |
|  | 2008 | 1148 | 18.8 | 8.3 | 5.2 | 34.5 | 33.2 | 0.0 | 26.0 | 7.2 | 18.8 | 17.9 |  | 19.4 |
| International Avg. | 1995 | 323 | 62.4 | 4.8 | 3.1 | 17.1 | 12.6 | 1.0 | 5.5 | 6.1 | 62.4 | 58.1 |  | 65.2 |
|  | 2008 | 1222 | 54.6 | 5.1 | 5.2 | 20.8 | 14.3 | 0.0 | 10.0 | 4.2 | 54.6 | 54.4 |  | 55.0 |

[^2]
### 8.2.2 Item-by-Country Interaction

Although countries are expected to exhibit some variation in performance across items, in general countries with high average performance on the assessment should perform relatively well on all of the items, and low-scoring countries should do less well on all of items. When this does not occur (i.e., when a high-performing country has a low performance on an item on which other countries are doing well), there is said to be an item-by-country interaction. When large, such item-by-country interactions may be a sign that an item is flawed in some way, and steps should be taken to address the problem.

To assist in detecting sizeable item-by-country interactions, the TIMSS \& PIRLS International Study Center produced a graphical display for each item showing the difference between each country's Rasch item difficulty and the average Rasch item difficulty across all countries. Exhibit 8.4 provides an example of a TIMSS Advanced 2008 item-by-country interaction display. The difference in Rasch item difficulty for each country is presented as a 95 percent confidence interval, which includes a built-in Bonferroni correction for multiple comparisons. The limits for this confidence interval were computed as follows:

$$
\begin{aligned}
& \text { Upper Limit }=\text { RDIFF }_{i .}-\text { RDIFF }_{i k}+S E\left(\text { RDIFF }_{i k}\right) \cdot Z_{b} \\
& \text { Lower Limit }=R D I F F_{i .}-\text { RDIFF }_{i k}-S E\left(\text { RDIFF }_{i k}\right) \cdot Z_{b}
\end{aligned}
$$

Where $R D I F F_{i k}$ was the Rasch difficulty of item $i$ in country $k, R D I F F_{i}$. was the average difficulty of item $i$ across all countries, $\operatorname{SE}\left(R D I F F_{i k}\right)$ was the standard error of the Rasch difficulty of item $i$ in country $k$, and $Z_{b}$ was the critical value from the $Z$ distribution corrected for multiple comparisons using the Bonferroni procedure.

## Exhibit 8.4 Sample Plot of Item-by-Country Interaction for a TIMSS Advanced 2008 Item

TIMSS Advanced 2008 - Plot of Item—by—Country Interactions
ItemName=W5_06 UniquelD=MA23157 Label=Find point of maxima and inflection


### 8.2.3 Trend Item Analysis

Because an important part of the TIMSS Advanced 2008 assessment was measuring trends across the 1995 and 2008 assessment cycles, an additional review step ensured that the trend items had similar characteristics in both cycles (i.e., an item that was relatively easy in 1995 should have been relatively easy in 2008). The comparison between cycles was made in a number of ways. For each trend country, almanacs of item statistics displayed the percentage of students within each score category (or response option for multiple-choice items) for each cycle, as well as the difficulty of the item and the percent correct by gender. While some changes were anticipated as countries' overall achievement may have improved or declined, items were noted if the difference between the Rasch difficulties across the two cycles for a particular country was greater than 2 logits.

The TIMSS \& PIRLS International Study Center used two different graphical displays to examine the differences item difficulties. The first of these, shown in Exhibit 8.5, displays the difference in Rasch item difficulties between 1995 and 2008. A positive difference indicated that an item was relatively easier in a country in 2008, and a negative difference indicated that an item was relatively more difficult. The second, Exhibit 8.6, shows the performance of a given country on all trend items simultaneously. For each country, the graph plotted the 1995 Rasch difficulty of every trend item in 1995 against its 2008 Rasch difficulty. Where there were no differences between the difficulties in 1995 and 2008, the data points aligned on or near the diagonal.

## Exhibit 8.5 Sample Plot of Difference in Rasch Item Difficulties for a Trend Item

TIMSS Advanced 2008 Trend - Plot of Difference in Rasch Difficulties
ItemName=M3_O6A UniquelD=MA13026A Label=Triangle abc/retlection


## Exhibit 8.6 Sample Plot of Rasch Item Difficulties across Trend Items by Country

TIMSS Advanced 2008 Trend - Plot of Rasch Difficulties by Country - M


### 8.2.4 Reliability

Gauging the reliability of the TIMSS Advanced 2008 assessments was a critical quality control step in reviewing the items. There were two aspects of reliability under review. The set of items selected as part of the advanced mathematics and physics assessments needed to constitute a cohesive whole measuring their respective domains, a quality known as test reliability. Also, the scoring of the constructedresponse items had to meet specific reliability criteria in terms of consistent within-country and cross-country scoring.

### 8.2.4.1 Test Reliability

Exhibit 8.7 displays the advanced mathematics and physics test reliability coefficients for every country. These coefficients are the median Cronbach's alpha reliability across the four test booklets of advanced mathematics and physics. In general, median reliabilities were relatively high in both subjects with the international median at 0.82 for advanced mathematics and 0.80 for physics. All median reliabilities were at least 0.70 , except for physics in Lebanon, where the median reliability was 0.68 .

Exhibit 8.7 Cronbach's Alpha Reliability Coefficients for TIMSS Advanced 2008

| Country | Reliability Coefficient |  |
| :--- | :---: | :---: |
|  | Advanced Mathematics | Physics |
| Armenia | 0.87 | 0.82 |
| Iran, Islamic Rep. of | 0.90 | 0.85 |
| Italy | 0.84 | 0.75 |
| Lebanon | 0.80 | 0.68 |
| Netherlands | 0.70 | 0.74 |
| Norway | 0.78 | 0.79 |
| Philippines | 0.79 | - |
| Russian Federation | 0.88 | 0.88 |
| Slovenia | 0.83 | 0.81 |
| Sweden | 0.80 | 0.80 |
| International Median | 0.82 | 0.80 |

### 8.2.4.2 Scoring Reliability for Constructed-response Items

About one third of the items in the TIMSS Advanced 2008 assessment were constructed-response items, comprising nearly half of the score points for the assessment. ${ }^{7}$ An essential requirement for use of such items is that they be reliably scored by all participants. That is, a particular student response should receive the same score, regardless of the scorer. In conducting TIMSS Advanced 2008, measures taken to ensure that the constructed-response items were scored reliably in all countries, and these measures included developing scoring guides for each constructed-response question (that provided descriptions of acceptable responses for each score point value) ${ }^{8}$ as well as providing extensive training in the application of the scoring guides.

## Within-Country Scoring Reliability

To gather and document information about the within-country agreement among scorers, a random sample of approximately 25 percent of the assessment booklets was selected to be scored independently by two scorers. The inter-scorer agreement for each item in each country was examined as part of the item review process. The average and range of the within-country percentage agreement across all items for both grades are presented in Exhibit 8.8 for both advanced mathematics and physics.

Scoring reliability was high on average across countries. The percent agreement on the correctness score across all countries was 98 percent in advanced mathematics and 97 percent in physics. All countries had an average percent agreement on the correctness score above 94 percent in advanced mathematics and above 91 percent in physics.

Exhibit 8.8 Within-country Scoring Reliability for TIMSS Advanced 2008 Constructed-Response Items

| Advanced Mathematics |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Correctness Score Agreement |  |  | Diagnostic Score Agreement |  |  |
| Countries | Average of Percent Agreement Across Items | Range of Percent Agreement |  | Average of Percent Agreement Across Items | Range of Percent Agreement |  |
|  |  | Min | Max |  | Min | Max |
| Armenia | 100 | 98 | 100 | 97 | 86 | 100 |
| Iran, Islamic Rep. of | 98 | 90 | 100 | 95 | 89 | 100 |
| Italy | 100 | 97 | 100 | 98 | 94 | 100 |
| Lebanon | 100 | 99 | 100 | 99 | 97 | 100 |
| Netherlands | 94 | 72 | 100 | 91 | 65 | 99 |
| Norway | 99 | 98 | 100 | 98 | 95 | 100 |
| Philippines | 98 | 93 | 100 | 95 | 85 | 100 |
| Russian Federation | 97 | 86 | 100 | 95 | 86 | 100 |
| Slovenia | 100 | 99 | 100 | 99 | 97 | 100 |
| Sweden | 97 | 88 | 100 | 93 | 83 | 99 |
| International Average | 98 | 92 | 100 | 96 | 88 | 100 |


| Physics |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Correctness Score Agreement |  |  | Diagnostic Score Agreement |  |  |
|  | Average of Percent Agreement Across Items | Range of Percent Agreement |  | Average of Percent Agreement Across Items | Range of Percent Agreement |  |
|  |  | Min | Max |  | Min | Max |
| Armenia | 99 | 93 | 100 | 97 | 93 | 100 |
| Iran, Islamic Rep. of | 96 | 91 | 100 | 90 | 71 | 99 |
| Italy | 99 | 94 | 100 | 97 | 86 | 100 |
| Lebanon | 99 | 93 | 100 | 98 | 92 | 100 |
| Netherlands | 91 | 80 | 99 | 85 | 71 | 97 |
| Norway | 97 | 90 | 100 | 94 | 87 | 100 |
| Russian Federation | 96 | 89 | 100 | 93 | 83 | 99 |
| Slovenia | 100 | 98 | 100 | 99 | 95 | 100 |
| Sweden | 97 | 89 | 99 | 93 | 83 | 99 |
| International Average | 97 | 91 | 100 | 94 | 85 | 99 |

## Cross-Country Scoring Reliability

Because of the different languages used by the countries participating in TIMSS Advanced 2008, establishing the reliability of constructedresponse scoring across all countries was not feasible. However, TIMSS Advanced 2008 did conduct a cross-country study of scoring using English as a common language. A sample of student responses from a pilot study carried out in English was provided to countries. It included 100 student responses to each of nine advanced mathematics items and nine physics items. This set of 1,800 student responses in English was then scored independently in each country that had two scorers proficient in English. In all, 14 scorers from 7 countries participated in the study. Scoring for this study took place shortly after the other scoring reliability activities were completed. Making all possible pairwise comparisons among scorers gave 91 comparisons for each student response to each item. This resulted in 9,100 total comparisons when aggregated across all 100 student responses to an item. Agreement across countries was defined in terms of the percentage of these comparisons that were in agreement.

Exhibit 8.9 shows that scorer reliability across countries was high for advanced mathematics, with the percent agreement averaging 94 percent across the nine items for the correctness score and 90 percent for the diagnostic score. For physics, the percent agreement averaged 88 percent across the nine items for the correctness score and 80 percent for the diagnostic score.

Exhibit 8.9 Cross-Country Scoring Reliability for TIMSS Advanced 2008

| Advanced Mathematics |  |  |  |
| :---: | :---: | :---: | :---: |
| Item Label | Total Valid Comparisons | Percent Agreement |  |
|  |  | Correctness Score Agreement | Diagnostic Score Agreement |
| M4_04-MA23201 | 9100 | 88 | 88 |
| M4_08-MA23043 | 9100 | 95 | 90 |
| M5_03-MA23054 | 9100 | 94 | 92 |
| M5_05 - MA23131A | 9100 | 99 | 99 |
| M5_05-MA23131B | 9100 | 98 | 97 |
| M5_10-MA23094 | 9100 | 89 | 81 |
| M6_07-MA23198 | 9100 | 94 | 92 |
| M7_03-MA23141 | 9100 | 98 | 83 |
| M7_11-MA23170 | 9100 | 90 | 84 |
| Average Percent Agreement |  | 94 | 90 |
| Physics |  |  |  |
| Item Label | Total Valid Comparisons | Percent Agreement |  |
|  |  | Correctness Score Agreement | Diagnostic Score Agreement |
| P4_07-PA23053 | 9100 | 86 | 82 |
| P4_09-PA23119 | 9100 | 80 | 67 |
| P4_11-PA23066 | 9100 | 83 | 81 |
| P5_03-PA23035 | 9100 | 95 | 85 |
| P5_05-PA23012 | 9100 | 88 | 74 |
| P5_07-PA23051 | 9100 | 94 | 87 |
| P6_05-PA23022 | 9100 | 83 | 76 |
| P7_05-PA23034 | 9100 | 97 | 95 |
| P7_06-PA23044 | 9100 | 90 | 70 |
| Average Percent Agre |  | 88 | 80 |

### 8.2.5 Summary of TIMSS Advanced 2008 Item Statistics Review

Based on the information from the comprehensive collection of item analyses and reliability data that were computed and summarized for TIMSS Advanced 2008, the TIMSS \& PIRLS International Study Center thoroughly reviewed all item statistics for every participating country to ensure that the items were performing comparably across countries. Specifically, items with the following problems were considered for possible deletion from the international database:

- An error was detected during the TIMSS Advanced 2008 translation verification but was not corrected before test administration.
- Data checking revealed a multiple-choice item with more or fewer options than in the international version.
- The item analysis showed an item to have a negative point-biserial, or, for an item with more than 1 score point, a non-monotonic relationship between score level and total score.
- The item-by-country interaction results showed a large negative interaction for a particular country.
- For constructed-response items, the within-country scoring reliability data showed a score agreement of less than 70 percent.
- For trend items, an item performed substantially differently in 2008 compared to 1995, or an item was not included in the 1995 assessment for a particular country.
When the item statistics indicated a problem with an item, the documentation from the translation verification ${ }^{9}$ was used as an aid in checking the test booklets. If a question remained about potential translation or cultural issues, however, then the National Research Coordinator was consulted before deciding how the item should be
treated. If a problem was detected by the TIMSS \& PIRLS International Study Center (such as a negative point-biserial for a correct answer or too few options for a multiple-choice item), the item was deleted from the international scaling.

The checking of the TIMSS Advanced 2008 achievement data involved a review of 143 items for 10 countries and resulted in the detection of very few items that were inappropriate for international comparisons. The few items singled out in the review process were mostly items with differences attributable to either translation or printing problems. The following is a list of deleted items as well as a list of recodings made to constructed-response items.

## Advanced Mathematics

## Items deleted

## ALL COUNTRIES

M2_O4 (MA13014) - attractive distractor

## LEBANON

M5_o8 (MA23082) - printing error

## Constructed-response items needing category recoding

## ALL COUNTRIES

M3_o6A (MA13026A) - recode 11 to 71
M3_o6B (MA13026B) - recode 11 to 72
M3_o8 (MA13028) - recode 20 to 10, 10 to 70, 11 to 71,12 to 72
M4_07 (MA23166) - recode 20 to 10, 21 to 11, 10 to 70,11 to 71,70 to 72

## Physics

## Items deleted

## ALL COUNTRIES

P1_07 (PA13007) - attractive distractor
P2_10 (PA13020) - low discrimination
P3_07B (PA13027B) - percent omitted too high

## SWEDEN

P4_04 (PA 23104) - negative discrimination

## Constructed-response items needing category recoding

## ALL COUNTRIES

P3_03 (PA13023) - recode 11 to 72
P3_05 (PA13025) - recode 29 to 19
P3_06 (PA13026) - recode 20 to 10, 29 to 19, 10 to 72 , 19 to 79
P3_07A (PA13027A) - recode 20 to 10, 21 to 11, 22 to 19, 29 to 19, 10 to 70,11 to 71
P4_10 (PA23088) - recode 11 to 71
P6_04 (PA23072) - recode 20 to 10 , 10 to 11
P6_07 (PA23078) - recode 20 to 10,10 to 11, 11 to 12

### 8.3 The TIMSS Advanced 2008 Scaling Methodology ${ }^{10}$

The IRT scaling approach used by TIMSS was developed originally by Educational Testing Service for use in the U.S. National Assessment of Educational Progress. It is based on psychometric models that were first used in the field of educational measurement in the 1950s and have become popular since the 1970s for use in large-scale surveys, test construction, and computer adaptive testing.

Three distinct IRT models, depending on item type and scoring procedure, were used in the analysis of the TIMSS Advanced assessment data. Each is a "latent variable" model that describes the probability that a student will respond in a specific way to an item in terms of the student's proficiency, which is an unobserved, or "latent," trait, and various characteristics (or "parameters") of the item. A threeparameter model was used with multiple-choice items, that were scored as correct or incorrect; and a two-parameter model, for constructedresponse items with two response options, that also were scored as correct or incorrect. Since each of these item types has two response categories, they are known as dichotomous items. A partial credit model was used with polytomous constructed-response items, i.e., those with more than two response options.

### 8.3.1 Proficiency Estimation Using Plausible Values

Most cognitive testing endeavors to assess the performance of individual students for the purposes of diagnosis, selection, or placement. Regardless of the measurement model used, whether classical test theory or item response theory, the accuracy of these measurements can be improved-that is, the amount of measurement error can be reduced-by increasing the number of items given to the individual. Thus, it is common to see achievement tests designed to provide information on individual students that contain more than 70 items. Since the uncertainty associated with estimates of individual student ability is negligible under these conditions, the distribution of student ability, or its joint distribution with other variables, can be approximated using each individual student's estimated ability.

For the distribution of proficiencies in large populations, more efficient estimates can be obtained from a matrix-sampling design such as that used in TIMSS Advanced. This design solicits relatively few
responses from each sampled student while maintaining a wide range of content representation when responses are aggregated across all students. With this approach, the advantage of estimating population characteristics more efficiently is offset to some degree by the inability to make precise statements about individuals. The uncertainty associated with individual student ability estimates becomes too large to be ignored.

Plausible values methodology was developed as a way to address this issue. Instead of first computing estimates of individual student abilities and then aggregating these to estimate population parameters, the plausible values approach uses all available data-students' responses to the items they were administered together with all background data-to estimate directly the characteristics of student populations and subpopulations. Although these directly estimated population characteristics could be used for reporting purposes, the usual plausible values approach generates multiple imputed scores, called plausible values, from the estimated ability distributions and uses these in analyses and reporting, making use of standard statistical software. By including all the available background data in the model, a process known as "conditioning," relationships between these background variables and the estimated proficiencies are appropriately accounted for in the plausible values. Because of this, analyses conducted using plausible values provide an accurate representation of these underlying relationships.

Plausible values are not intended to be estimates of individual student scores, but rather are imputed scores for similar studentsstudents with similar response patterns and background characteristics in the sampled population-that may be used to estimate population characteristics correctly. When the underlying model is correctly specified, plausible values provide consistent estimates of population
characteristics, even though they are not generally unbiased estimates of the proficiencies of the individuals with whom they are associated. Taking the average of the plausible values does not yield suitable estimates of individual student scores. ${ }^{11}$

### 8.4 Implementing the Scaling Procedures for the TIMSS Advanced Assessment Data

The application of IRT scaling and plausible values methodology to the data from the TIMSS Advanced 2008 assessments involved four major tasks: calibrating the achievement test items (estimating model parameters for each item), creating principal components from the student questionnaire data for use in conditioning, generating proficiency scores for advanced mathematics and for physics, and placing these proficiency scores on the scales-one for advanced mathematics and one for physics-used to report the results from the TIMSS Advanced assessments in 1995.

Before scaling the 2008 assessment data, however, the data from the 1995 assessments had to be rescaled from the one-parameter Rasch model used in 1995 to the multi-parameter models that have been in use in TIMSS since 1999.

### 8.4.1 Rescaling the Data from the TIMSS Advanced 1995 Assessments

The students' responses to the achievement items and to the questions in the student background questionnaire from the TIMSS Advanced 1995 international database provided the data for rescaling the TIMSS Advanced 1995 data. The TIMSS Advanced 1995 assessments included 68 items for advanced mathematics and 66 items for physics. These items were classified into the content and cognitive domains defined in the TIMSS Advanced 2008 Assessment Frameworks (Garden, et al., 2006) in preparation for trend scaling. Of the 134 items, 10 advanced
mathematics items and 4 physics items did not fit any framework classification and thus were omitted from the TIMSS Advanced 1995 rescaling since they were no longer appropriate for the domains specified in the 2008 frameworks. Also, one advanced mathematics item and three physics items were omitted for the reasons given in Section 8.2.5. Finally, one physics item released after the 1995 assessment (PA13052), was omitted because of poor discrimination. Some trend items that required recoding in the 2008 assessments also were recoded in the 1995 database.

All countries that participated in TIMSS Advanced 1995 were included in the item calibrations. Exhibit 8.10 presents the sample sizes for the countries included in the TIMSS Advanced 1995 item calibrations. ${ }^{12}$

Exhibit 8.10 Sample Sizes for Item Calibrations of the TIMSS Advanced 1995 Assessments

| Country | Advanced Mathematics | Physics |
| :--- | ---: | ---: |
| Australia | 645 | 661 |
| Austria | 782 | 777 |
| Canada | 2,781 | 2,367 |
| Cyprus | 391 | 368 |
| Czech Republic | 1,101 | 1,087 |
| France | 1,071 | 1,110 |
| Germany | 2,296 | 723 |
| Greece | 456 | 459 |
| Italy | 398 | - |
| Latvia | - | 708 |
| Lithuania | 734 | - |
| Norway | - | 1,048 |
| Russian Federation | $\mathbf{1 , 6 3 8}$ | 1,233 |
| Slovenia | 1,536 | 747 |
| Sweden | 1,001 | 1,012 |
| Switzerland | 1,404 | 1,371 |
| United States | $\mathbf{2 , 7 8 5}$ | 3,114 |
| Total | $\mathbf{1 9 , 0 1 9}$ | $\mathbf{1 6 , 7 8 5}$ |

[^3]The item calibrations were conducted separately for each subject by the TIMSS \& PIRLS International Study Center using the commercially-available Parscale software (Muraki \& Bock, 1991; version 4.1). The two- and three-parameter and polytomous IRT models were fitted to the data to produce item parameter estimates. These new estimated values can be found in Exhibits D. 1 and D. 2 of Appendix D. These parameter estimates then became part of the input for producing proficiency scores.

A principal components analysis was run separately within each country to generate input for the conditioning step. The estimated proficiency scores are conditioned on the student background variables to improve the reliability of sub-population reporting. Principal components analysis is used to reduce the number of conditioning variables to a manageable size. The usual TIMSS approach retains the number of principal components that account for at least 90 percent of the variability in the student background data. Since most countries in 1995 had small sample sizes, the 90 percent criterion was reduced to 70 percent to minimize over-specification in the conditioning model, provided the number of components retained did not exceed 10 percent of the sample size-in which case the number of components was limited to 10 percent of the sample size. Exhibit 8.11 displays the total number of variables considered for conditioning and the number of principal components selected for each country.

The generation of IRT proficiency scores was conducted separately for each country and for each subject using Educational Testing Service's MGROUP program (Sheehan, 1985; version 3.2). ${ }^{13}$ MGROUP takes as input the students' responses to the items they were given, the item parameters estimated at the calibration stage, and the conditioning variables, and generates as output the plausible values

Exhibit 8.11 Number of Variables and Principal Components for Conditioning in Rescaling the TIMSS Advanced 1995 Assessments

| Countries | Advanced Mathematics |  |  | Physics |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Number of Primary Conditioning Variables | Total Number of Principal Components | Number of Principal Components Retained | Number of Primary Conditioning Variables | Total <br> Number of Principal Components | Number of Principal Components Retained |
| Australia | 2 | 553 | 64 | 2 | 560 | 66 |
| Austria | 2 | 542 | 78 | 2 | 545 | 77 |
| Canada | 3 | 573 | 148 | 3 | 570 | 134 |
| Cyprus | 2 | 562 | 39 | 2 | 565 | 36 |
| Czech Republic | 2 | 565 | 110 | 2 | 549 | 108 |
| Denmark | 2 | 545 | 123 | 2 | 533 | 65 |
| France | 2 | 434 | 107 | 2 | 435 | 111 |
| Germany | 2 | 542 | 147 | 2 | 526 | 72 |
| Greece | 2 | 551 | 45 | 2 | 550 | 45 |
| Israel | 2 | 589 | 95 | 2 | 583 | 85 |
| Italy | 2 | 531 | 39 | - | - | - |
| Latvia | - | - | - | 2 | 579 | 70 |
| Lithuania | 2 | 542 | 73 | - | - | - |
| Norway | - | - | - | 3 | 579 | 104 |
| Russian Federation | 2 | 602 | 152 | 2 | 598 | 123 |
| Slovenia | 2 | 582 | 141 | 2 | 573 | 74 |
| Sweden | 2 | 576 | 100 | 2 | 571 | 101 |
| Switzerland | 4 | 544 | 127 | 4 | 546 | 125 |
| United States | 2 | 612 | 154 | 2 | 618 | 166 |

that represent student proficiency. Exhibit 8.12 shows the sample sizes of the countries for which proficiency scores were generated. ${ }^{14}$

The reporting metrics for the rescaled 1995 data were established to give the distribution of TIMSS Advanced 1995 proficiency scores in advanced mathematics and in physics a mean of 500 and a standard deviation of 100, with all 1995 countries included in the item calibrations contributing equally. Extreme scale values were truncated, giving plausible values a minimum of 5 and a maximum of 995 .

[^4]
## Exhibit 8.12 Sample Sizes for Proficiency Estimation of the TIMSS Advanced 1995 Assessments

| Country | Advanced Mathematics | Physics |
| :--- | ---: | ---: |
| Australia | 645 | 661 |
| Austria | 782 | 777 |
| Canada | 2,781 | 2,367 |
| Cyprus | 391 | 368 |
| Czech Republic | 1,101 | 1,087 |
| Denmark | 1,388 | 654 |
| France | 1,071 | 1,110 |
| Germany | 2,296 | 723 |
| Greece | 456 | 459 |
| Israel | 953 | 853 |
| Italy | 398 | - |
| Latvia | - | 708 |
| Lithuania | 734 | - |
| Norway | - | 1,048 |
| Russian Federation | 1,638 | 1,233 |
| Slovenia | 1,536 | 747 |
| Sweden | 1,001 | 1,012 |
| Switzerland | 1,404 | 1,371 |
| United States | 2,785 | 3,114 |
| Total | $\mathbf{2 1 , 3 6 0}$ | $\mathbf{1 8 , 2 9 2}$ |

### 8.4.2 Calibrating the TIMSS Advanced 2008 Assessment Data

As described in the TIMSS Advanced 2008 Assessment Frameworks (Garden, et al., 2006), the TIMSS Advanced 2008 assessments consisted of a total of seven advanced mathematics blocks and seven physics blocks, distributed across eight assessment booklets. Each block contained either advanced mathematics or physics items, drawn from a range of content and cognitive domains. The seven mathematics blocks were designated $\mathrm{M}_{1}$ through $\mathrm{M}_{7}$, and the seven physics blocks $P_{1}$ through $P_{7}$. Blocks $M_{1}$ through $M_{3}$ and $P_{1}$ through $P_{3}$ contained
items that were previously used in the 1995 assessments, whereas blocks $\mathrm{M}_{4}$ through $\mathrm{M}_{7}$ and $\mathrm{P}_{4}$ through $\mathrm{P}_{7}$ consisted of newly-developed items for the 2008 assessments. Each assessment booklet contained three blocks of either all advanced mathematics items or all physics items. The booklets were distributed among the students in each sampled class according to a scheme that ensured equivalent random samples of students responding to each booklet.

Separate IRT scales were constructed for reporting overall student achievement in advanced mathematics and in physics. Concurrent item calibrations were conducted by the TIMSS \& PIRLS International Study Center using Parscale, and included data from the TIMSS Advanced 2008 assessments and the TIMSS Advanced 1995 assessments to measure trends from 1995. The calibrations used all available data from each country's TIMSS Advanced student samples, which were weighted such that each country contributed equally.

The first step in constructing the scales for TIMSS Advanced 2008 was to estimate the IRT model item parameters for each item on each of the scales through a concurrent calibration of both sets of assessment data-1995 and 2008. It was then possible to obtain the mean and standard deviation of the latent ability distributions of students in both assessments using item paramters from the concurrent calibration. The difference between these two distributions was the change in achievement from 1995 to 2008. The second step was to find the linear transformation that transformed the distribution of the 1995 assessment data under the 1995-2008 concurrent calibration to match the distribution of these data under the 1995 calibration. The third step was to apply this same transformation to the 2008 assessment data scaled using the concurrent calibration. This placed the 2008 assessment data on the metric of the 1995 assessment-i.e., a scale with a mean of 500 and a standard deviation of 100 .

Exhibit 8.13 illustrates how the concurrent calibration approach was applied in the context of TIMSS Advanced 2008 trend scaling. The observed gap between the distribution of the 1995 data under the 1995 item calibration and the 1995 data under the 1995-2008 concurrent calibration was small and arose from slight differences in the item parameter estimations, which in turn were due mostly to the 1995 assessment data being calibrated with the 2008 assessment data. The linear transformation removed this gap by shifting the two distributions from the concurrent calibration such that the distribution of the 1995 assessment data from the concurrent calibration aligned with the distribution of the 1995 assessment data from the 1995 calibration, while preserving the gap between the 1995 and 2008 assessment data under the concurrent calibration. This latter gap was the change in achievement between the previous and current assessments that TIMSS Advanced set out to measure as its trend.

## Exhibit 8.13 Concurrent Calibration Model Used for TIMSS Advanced 2008




Having estimated the item parameters from the 1995-2008 concurrent calibration, new achievement distributions were generated by applying these item parameters to the 1995 assessment and to the 2008 assessment data. Following the procedure outlined above, the next step was to identify the linear transformation that transformed the 1995 assessment distribution generated by the 1995-2008 concurrent calibration item parameters to match the 1995 assessment distribution generated by the item parameters from the 1995 rescaling, and to apply this same transformation to the 2008 assessment data distribution (also generated by the concurrent calibration item parameters).

Exhibit 8.14 shows the distribution of items included in the TIMSS Advanced 2008 concurrent calibrations for reporting trends in overall advanced mathematics and physics. All the data from both the 1995 and 2008 assessments were included. Items were categorized as unique to the 1995 assessment, common to both assessments, or unique to the 2008 assessment. For advanced mathematics, the 2008 assessment contributed 45 items worth 51 score points that were unique to 2008 and 26 items worth 28 score points that also were included in the 1995 assessment. The 1995 assessment also contributed 31 items worth 40 score points that were released in 1995 . For physics, the 2008 assessment contributed 45 items worth 51 score points that were unique to 2008 and 23 items worth 26 score points that also were included in the 1995 assessment. The 1995 assessment also contributed 35 items worth 42 score points that were released in 1995.

Exhibit 8.14 Items Included in the TIMSS Advanced 2008 Concurrent Item Calibrations

| TIMSS 2008 Trend Scales | Items Unique to the TIMSS Advanced 2008 Assessments |  | Items Common to the TIMSS Advanced 2008 and 1995 Assessments |  | Items Unique to the TIMSS Advanced 1995 Assessments |  | TOTAL |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Number | Points | Number | Points | Number | Points | Number | Points |
| Advanced Mathematics | 45 | 51 | 26 | 28 | 31 | 40 | 102 | 119 |
| Physics | 45 | 51 | 23 | 26 | 35 | 42 | 103 | 119 |

Because of the small number of countries that participated in both TIMSS Advanced assessments, concurrent item calibrations were conducted using data from all the countries that participated in either the 1995 assessments or the 2008 assessments. To construct the advanced mathematics scale, the calibration included 19,019 students from 15 countries in the 1995 assessment and 22,242 students from 10 countries in the 2008 assessment. The item parameters established in this calibration were used subsequently for estimating student proficiency scores in advanced mathematics for the 10 countries that participated in TIMSS Advanced 2008. The national samples included in the calibration for reporting trends in advanced mathematics are presented in Exhibit 8.15.

Similarly, to construct the physics scale, the item calibration was conducted using data from countries that participated in either the 1995 assessment or the 2008 assessment. The physics concurrent calibration included 16,785 students from 15 countries in the 1995 assessment and 16,489 students from 9 countries in the 2008 assessment. The item parameters obtained in this calibration were used subsequently for estimating student proficiency scores in physics for the nine countries that participated in TIMSS Advanced 2008. Exhibit 8.16 presents the national samples included in the calibration for reporting trends in Physics.

Exhibits D. 3 and D. 4 of Appendix D display the item parameters for advanced mathematics and physics, respectively, generated from the concurrent calibration of the 1995 and the 2008 data. As a by-product of the calibrations, interim scores in advanced mathematics and physics were produced for use in constructing conditioning variables.

Exhibit 8.15 Sample Sizes for Concurrent Item Calibration of Advanced Mathematics for the TIMSS Advanced 1995 and 2008 Assessments

| Country | 1995 Assessment | 2008 Assessment |
| :---: | :---: | :---: |
| Countries in Both Cycles |  |  |
| Italy | 398 | 2,143 |
| Russian Federation | 1,638 | 3,185 |
| Slovenia | 1,536 | 2,156 |
| Sweden | 1,001 | 2,303 |
| Countries in 1995 |  |  |
| Australia | 645 | - |
| Austria | 782 | - |
| Canada | 2,781 | - |
| Cyprus | 391 | - |
| Czech Republic | 1,101 | - |
| France | 1,071 | - |
| Germany | 2,296 | - |
| Greece | 456 | - |
| Lithuania | 734 | - |
| Switzerland | 1,404 | - |
| United States | 2,785 | - |
| Countries in 2008 |  |  |
| Armenia | - | 858 |
| Iran, Islamic Rep. of | - | 2,425 |
| Lebanon | - | 1,612 |
| Netherlands | - | 1,537 |
| Norway | - | 1,932 |
| Philippines | - | 4,091 |
| Total | 19,019 | 22,242 |

$\begin{array}{ll}\text { Exhibit 8.16 } & \text { Sample Sizes for Concurrent Item Calibration of Physics for the } \\ & \text { TIMSS Advanced } 1995 \text { and } 2008 \text { Assessments }\end{array}$


| Norway | 1,048 | 1,640 |
| :--- | ---: | ---: |
| Russian Federation | 1,233 | 3,166 |
| Slovenia | 747 | 1,097 |
| Sweden | 1,012 | 2,291 |


| Countries in 1995 |  |  |
| :--- | ---: | :---: |
| Australia | 661 | - |
| Austria | 777 | - |
| Canada | 3667 | - |
| Cyprus | 1,087 | - |
| Czech Republic | 1,110 | - |
| France | 723 | - |
| Germany | 459 | - |
| Greece | 708 | - |
| Latvia | 1,371 | - |
| Switzerland | 3,114 | - |
| United States |  |  |
| Countries in 2008 |  |  |


| Armenia | - | 894 |
| :--- | :---: | ---: |
| Iran, Islamic Rep. of | - | 2,434 |
| Italy | - | 1,861 |
| Lebanon | - | $\mathbf{1 , 5 9 5}$ |
| Netherlands | - | $\mathbf{1 , 5 1 1}$ |
| Total | $\mathbf{1 6 , 7 8 5}$ | $\mathbf{1 6 , 4 8 9}$ |

### 8.4.3 Omitted and Not-Reached Responses

Apart from missing data on items that by design were not administered to a student, missing data could also occur because a student did not answer an item-whether because the student did not know the answer, omitted it by mistake, or did not have time to attempt the item. An item was considered not reached when the item itself and the item
immediately preceding it were not answered, and there were no other items completed in the remainder of the booklet.

In TIMSS Advanced 2008, as in TIMSS Advanced 1995 and previous TIMSS assessments, not-reached items were treated differently in estimating item parameters than they were in generating student proficiency scores. In estimating the values of the item parameters, items in the TIMSS Advanced assessment booklets that were considered not to have been reached by students were treated as if they had not been administered. This approach was considered optimal for parameter estimation. However, not-reached items were always considered as incorrect responses when student proficiency scores were generated.

### 8.4.4 Evaluating the Fit of IRT Models to the TIMSS Advanced 2008 Data

After the concurrent item calibrations were completed, checks were performed to verify that the item parameters obtained from Parscale adequately reproduced the observed distribution of student responses across the proficiency continuum. The fit of the IRT models to the TIMSS Advanced data was examined by comparing the item response function curves generated using the item parameters estimated from the data with the empirical item response functions calculated from the posterior distributions of the proficiencies for each student that responded to an item. When the empirical results fall near the fitted curves for any given item, the IRT model fits the data well and leads to more accurate and reliable measurement of the underlying proficiency scale. Graphical plots of these response function curves are called item characteristic curves (ICC).

Exhibit 8.17 shows an ICC of the empirical and fitted item response functions for a dichotomous multiple-choice item. In the graph, the horizontal axis represents the proficiency scale; and the vertical axis,
the probability of a correct response. The fitted curve based on the estimated item parameters is shown as a solid line. Empirical results are represented by circles. The empirical results were obtained by dividing the proficiency scale into intervals of equal size and then counting the number of students responding to the item whose estimated a-priori (EAP) scores from Parscale fell in each interval. Then the proportion of students in each interval that responded correctly to the item was calculated. In the exhibit, the center of each circle represents this empirical proportion of correct responses. The size of each circle is proportional to the number of students contributing to the estimation of its empirical proportion correct.

Exhibit 8.17 Example Item Response Function for a TIMSS Advanced 2008 Dichotomous Item


Exhibit 8.18 contains an ICC of the empirical and fitted item response functions for a polytomous constructed-response item with three response categories- 0,1 , and 2 points. As for the dichotomous item plot, the horizontal axis represents the proficiency scale, but the vertical axis represents the probability of having a response in a given response category. The fitted curves based on the estimated item parameters are shown as solid lines. Empirical results are represented by circles. The interpretation of the circles is the same as in Exhibit 8.17. The curve starting at the top left of the chart plots the probability of a score of zero on the item, which decreases as proficiency increases.

Exhibit 8.18 Example Item Response Function for a TIMSS Advanced 2008 Polytomous Item
Probability of a Correct Response for Estimated Proficiency - Mathematics - CLB

Item=MA23157 BlkSeq=M5_06 Ncat=3 $a=0.730 \quad b=0.80283$ step $1=0.564$ step2 $=-0.564$


The bell-shaped curve shows the probability of a score of 1 pointstarting low for low-ability students, reaching a maximum for mediumability students, and decreasing for high-ability students. The curve ending at the top right corner of the chart shows the probability of a score of 2 points-full credit, starting low for low-ability students and increasing as proficiency increases.

### 8.4.5 Variables for Conditioning the TIMSS Advanced 2008 Data

Because there were so many background variables that could be used in conditioning, TIMSS Advanced followed the practice established by NAEP and followed by other large-scale studies of using principal components analysis to reduce the number of variables while explaining most of their common variance. Principal components for the TIMSS Advanced background data were constructed as follows:

- For categorical variables (questions with a small number of fixed response options), a "dummy coded" variable was created for each response option, with a value of 1 if the option was chosen and zero otherwise. If a student omitted or was not administered a particular question, all dummy coded variables associated with that question were assigned the value zero.
- Background variables with numerous categories (such as year of birth or time spent doing homework) were recoded using criterion scaling. ${ }^{15}$ This was done by replacing each response option with an interim achievement score. For the overall advanced mathematics scale, the interim achievement score was the advanced mathematics score produced from the item calibration. For the overall physics scale, the interim achievement score was the physics score produced from the item calibration.
- Separately for each TIMSS country, all the dummy-coded and criterion-scaled variables were included in a principal components
analysis. Those principal components accounting for 90 percent of the variance of the background variables were retained for use as conditioning variables. However, if the selected number of principal components exceeded 5 percent of the student sample size, the number of selected principal components was reduced to 5 percent of the student sample size. Because the principal components analysis was performed separately for each country, different numbers of principal components were required to account for 90 percent of the common variance in each country's background variables.

In addition to the principal components, student gender (dummy coded), the language of the test (dummy coded), and an indicator of the class in the school to which the student belonged (criterion scaled) were included as primary conditioning variables, thereby accounting for most of the variance among students and preserving the betweenand within-class variance structure in the scaling model. Conditioning variables were needed for all the TIMSS Advanced 2008 participants, as well as for all the TIMSS Advanced 1995 countries. Exhibits 8.19 and 8.20 show the total number of variables that were considered for conditioning and the number of principal components selected for each country for advanced mathematics and physics, respectively.

Exhibit 8.19 Number of Variables and Principal Components for Conditioning Advanced Mathematics in TIMSS Advanced 2008

| Countries | 1995 Assessment |  |  | 2008 Assessment |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Number of Primary Conditioning Variables | Total <br> Number of Principal Components | Number of Principal Components Retained | Number of Primary Conditioning Variables | Total <br> Number of Principal Components | Number of Principal Components Retained |
| Armenia | - | - | - | 2 | 271 | 42 |
| Australia | 2 | 553 | 64 | - | - | - |
| Austria | 2 | 542 | 78 | - | - | - |
| Canada | 3 | 573 | 148 | - | - | - |
| Cyprus | 2 | 562 | 39 | - | - | - |
| Czech Republic | 2 | 565 | 110 | - | - | - |
| Denmark | 2 | 545 | 123 | - | - | - |
| France | 2 | 434 | 107 | - | - | - |
| Germany | 2 | 542 | 147 | - | - | - |
| Greece | 2 | 551 | 45 | - | - | - |
| Iran, Islamic Rep. of | - | - | - | 2 | 279 | 121 |
| Israel | 2 | 589 | 95 | - | - | - |
| Italy | 2 | 531 | 39 | 2 | 270 | 107 |
| Lebanon | - | - | - | 3 | 277 | 80 |
| Lithuania | 2 | 542 | 73 | - | - | - |
| Netherlands | - | - | - | 2 | 267 | 76 |
| Norway | - | - | - | 2 | 270 | 96 |
| Philippines | - | - | - | 2 | 276 | 156 |
| Russian Federation | 2 | 602 | 152 | 2 | 277 | 157 |
| Slovenia | 2 | 582 | 141 | 2 | 270 | 107 |
| Sweden | 2 | 576 | 100 | 2 | 268 | 115 |
| Switzerland | 4 | 544 | 127 | - | - | - |
| United States | 2 | 612 | 154 | - | - | - |

Exhibit 8.20 Number of Variables and Principal Components for Conditioning Physics in TIMSS Advanced 2008

| Countries | 1995 Assessment |  |  | 2008 Assessment |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Number of Primary Conditioning Variables | Total <br> Number of Principal Components | Number of Principal Components Retained | Number of Primary Conditioning Variables | Total <br> Number of Principal Components | Number of Principal Components Retained |
| Armenia | - | - | - | 2 | 275 | 44 |
| Australia | 2 | 560 | 66 | - | - | - |
| Austria | 2 | 545 | 77 | - | - | - |
| Canada | 3 | 570 | 134 | - | - | - |
| Cyprus | 2 | 565 | 36 | - | - | - |
| Czech Republic | 2 | 549 | 108 | - | - | - |
| Denmark | 2 | 533 | 65 | - | - | - |
| France | 2 | 435 | 111 | - | - | - |
| Germany | 2 | 526 | 72 | - | - | - |
| Greece | 2 | 550 | 45 | - | - | - |
| Iran, Islamic Rep. of | - | - | - | 2 | 282 | 121 |
| Israel | 2 | 583 | 85 | - | - | - |
| Italy | - | - | - | 2 | 205 | 93 |
| Lebanon | - | - | - | 3 | 281 | 79 |
| Latvia | 2 | 579 | 70 | - | - | - |
| Netherlands | - | - | - | 2 | 272 | 75 |
| Norway | 3 | 579 | 104 | 2 | 270 | 82 |
| Russian Federation | 2 | 598 | 123 | 2 | 283 | 158 |
| Slovenia | 2 | 573 | 74 | 2 | 272 | 54 |
| Sweden | 2 | 571 | 101 | 2 | 268 | 114 |
| Switzerland | 4 | 546 | 125 | - | - | - |
| United States | 2 | 618 | 166 | - | - | - |

### 8.4.6 Generating IRT Proficiency Scores for the TIMSS Advanced 2008 Data

MGROUP was used to generate the IRT proficiency scores. Exhibit 8.21 shows the student sample sizes-from the 1995 assessments and the 2008 assessments-for which proficiency scores, using the item parameters obtained from the concurrent calibration, were generated on the overall advanced mathematics and physics scales.

Exhibit 8.21 Sample Sizes for TIMSS Advanced Proficiency Estimation

| Country | Advanced Mathematics |  | Physics |  |
| :---: | :---: | :---: | :---: | :---: |
|  | 1995 Assessment | 2008 Assessment | 1995 Assessment | 2008 Assessment |
| Armenia | - | 858 | - | 894 |
| Australia | 645 | - | 661 | - |
| Austria | 782 | - | 777 | - |
| Canada | 2,781 | - | 2,367 | - |
| Cyprus | 391 | - | 368 | - |
| Czech Republic | 1,101 | - | 1,087 | - |
| France | 1,071 | - | 1,110 | - |
| Germany | 2,296 | - | 723 | - |
| Greece | 456 | - | 459 | - |
| Iran, Islamic Rep. of | - | 2,425 | - | 2,434 |
| Italy | 398 | 2,143 | - | 1,861 |
| Latvia | - | - | 708 | - |
| Lebanon | - | 1,612 | - | 1,595 |
| Lithuania | 734 | - | - | - |
| Netherlands | - | 1,537 | - | 1,511 |
| Norway | - | 1,932 | 1,048 | 1,640 |
| Philippines | - | 4,091 | - | - |
| Russian Federation | 1,638 | 3,185 | 1,233 | 3,166 |
| Slovenia | 1,536 | 2,156 | 747 | 1,097 |
| Sweden | 1,001 | 2,303 | 1,012 | 2,291 |
| Switzerland | 1,404 | - | 1,371 | - |
| United States | 2,785 | - | 3,114 | - |
| Total | 19,019 | 22,242 | 16,785 | 16,489 |

### 8.4.7 Transforming the Advanced Mathematics and Physics Scores to Measure Trends

As part of rescaling the data from the TIMSS Advanced 1995 assessments using 2- and 3-parameter models as described in Section 8.4.1, the TIMSS Advanced reporting scales were established by setting the average of the mean scores of the countries included in the rescaling item calibrations to 500 and the standard deviation to 100 . To provide results for the 2008 assessments that would be directly comparable to the results from the 1995 assessments, the 2008 proficiency scores (plausible values) for advanced mathematics and physics had to be transformed to the TIMSS Advanced scales established with the 1995 data. This was accomplished through a linear transformation of the proficiency scores from the 1995-2008 concurrent calibration such that the 1995 proficiency distribution from the concurrent calibration aligned itself with the 1995 proficiency distribution from the 1995 rescaling calibration.

The means and standard deviations of the 1995 advanced mathematics and physics scores produced in 2008-the plausible values from the TIMSS Advanced 1995 assessment data based on the 1995-2008 concurrent item calibrations-were made to match the means and standard deviations of the scores calculated for the TIMSS Advanced 1995 assessment-the plausible values produced using the item calibration from scaling the 1995 assessment data-by applying the appropriate linear transformations. These linear transformations were given by:

$$
P V_{k, i}{ }^{*}=A_{k, i}+B_{k, i} \cdot P V_{k, i}
$$

where
$P V_{k, i}$ was plausible value $i$ for scale $k$ prior to transformation,
$P V_{k, i}{ }^{*}$ was plausible value $i$ for scale $k$ after transformation,
and $A_{k, i}$ and $B_{k, i}$ were the linear transformation constants for plausible value $i$ of scale $k$.

The linear transformation constants were obtained by first computing the international means and standard deviations of the proficiency scores for the overall advanced mathematics and physics scales using the plausible values from the 1995 scaling for the 1995 countries included in the concurrent calibration. Next, the same calculation was done using the plausible values from the 1995 assessment data based on the 1995-2008 concurrent calibration for the same set of countries. The linear transformation constants were thus defined as:

$$
\begin{aligned}
B_{k, i} & =\sigma_{k, i}^{*} / \sigma_{k, i} \\
A_{k, i} & =\mu_{k, i}^{*}-B_{k, i} \mu_{k, i}
\end{aligned}
$$

where
$\mu_{k, i}{ }^{*} \quad$ was the international mean of scale $k$ based on plausible value $i$ obtained from scaling the 1995 assessment data;
$\mu_{k, i} \quad$ was the international mean of scale $k$ based on plausible value $i$ from the TIMSS Advanced 1995 assessment data based on the 1995-2008 concurrent item calibrations;
$\sigma_{k, i}{ }^{*} \quad$ was the international standard deviation of scale $k$ based on plausible value $i$ obtained from scaling the 1995 assessment data;
$\sigma_{k, i} \quad$ was the international standard deviation of scale $k$ based on plausible value $i$ from the TIMSS Advanced 1995 assessment data based on the 1995-2008 concurrent item calibrations.

Exhibit 8.22 shows the linear transformation constants that were computed for TIMSS Advanced 2008 for the advanced mathematics and physics scales. Once these linear transformation constants had been established, all of the 2008 advanced mathematics and physics
proficiency scores-the plausible values generated from the 2008 assessment data-for all the participating countries were transformed by applying the linear transformations. This provided advanced mathematics and physics student achievement scores for the 2008 assessments that were directly comparable to the rescaled scores from the 1995 assessment data.

Exhibit 8.22 Linear Transformation Constants Applied to the TIMSS Advanced 2008 Proficiency Scores

| Scale | Plausible Value | TIMSS Advanced 1995 Scores Using 1995 Item Calibrations |  | TIMSS Advanced 1995 Scores <br> Using 1995-2008 <br> Concurrent Calibrations |  | $A_{k, i}$ | $B_{k, i}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Mean | Standard Deviation | Mean | Standard <br> Deviation |  |  |
| Advanced Mathematics | PV1 | 500.00109 | 99.99451 | 0.11465 | 0.88983 | 487.11733 | 112.37442 |
|  | PV2 | 500.00010 | 99.99951 | 0.10189 | 0.89226 | 488.58124 | 112.07392 |
|  | PV3 | 500.00047 | 99.99754 | 0.11694 | 0.89312 | 486.90788 | 111.96396 |
|  | PV4 | 500.00004 | 99.99981 | 0.10578 | 0.88989 | 488.11346 | 112.37280 |
|  | PV5 | 500.00107 | 99.99447 | 0.10317 | 0.89075 | 488.41976 | 112.25904 |
| Physics | PV1 | 500.00000 | 100.00000 | -0.02718 | 0.97905 | 502.77604 | 102.14010 |
|  | PV2 | 500.00000 | 100.00000 | -0.02970 | 0.98882 | 503.00383 | 101.13043 |
|  | PV3 | 500.00000 | 100.00000 | -0.02392 | 0.97793 | 502.44621 | 102.25631 |
|  | PV4 | 500.00000 | 100.00000 | -0.02718 | 0.98387 | 502.76209 | 101.63923 |
|  | PV5 | 500.00000 | 100.00000 | -0.03582 | 0.97809 | 503.66228 | 102.23961 |

Note: The means and standard deviations for advanced mathematics based on the 1995 item calibrations are affected by rare cases of very low scores that were truncated.

### 8.5 The TIMSS Advanced 2008 International Benchmarks of Student Achievement in Advanced Mathematics and Physics

To describe student performance at various points along the TIMSS 2008 advanced mathematics and physics achievement scales, TIMSS Advanced 2008 used scale anchoring to summarize and describe student achievement at three points on the advanced mathematics and physics scales-Advanced International Benchmark (625), High

International Benchmark (550), and Intermediate International Benchmark (475). For a description of performance at the international benchmarks, please see the TIMSS Advanced 2008 International Report (Mullis, Martin, Robitaille, \& Foy, 2009).

In brief, scale anchoring involves selecting benchmarks (scale points) on the TIMSS achievement scales to be described in terms of student performance and then identifying items that students scoring at those anchor points (the international benchmarks) can answer correctly. The items so identified were grouped by content domain within benchmarks for review by mathematics and physics experts. The committee members ${ }^{16}$ examined the content of each item and determined the kind of mathematics or physics knowledge or skill demonstrated by students who responded correctly to the item. They then summarized the detailed list of item competencies in a brief description of achievement at each international benchmark. This procedure resulted in a content-referenced interpretation of the achievement results that can be considered in light of the TIMSS Advanced 2008 advanced mathematics and physics frameworks.

As the first step, students scoring within 20 scale score points of each benchmark were identified for the benchmark analysis. The score ranges around each international benchmark and the number of students scoring in each range for advanced mathematics and physics are shown in Exhibit 8.23. The range of 20 points above and below a benchmark provided an adequate sample in each group, yet was small enough so that performance at each benchmark anchor point was still distinguishable from the next.

## Exhibit 8.23 Range Around Each International Benchmark and Number of Students Within Each Range

|  | Intermediate (475) | High (550) | Advanced (625) |
| :--- | ---: | ---: | ---: |
| Range of Scale Scores | $455-495$ | $530-570$ | $605-645$ |
| Advanced Mathematics | 2,826 | 2,752 | 1,138 |
| Physics | 2,201 | 2,369 | 1,327 |

Having identified the number of students scoring within each benchmark range, the next step was conducting the data analysis to determine which items anchored at each of the international benchmarks. An important feature of the scale anchoring method is that it yields descriptions of the performance demonstrated by students reaching each of the international benchmarks on the scales, and that the descriptions reflect demonstrably different accomplishments by students reaching each successively higher benchmark. Because the process entails the delineation of sets of items that students at each international benchmark are likely to answer correctly and that discriminate between one benchmark and the next, the criteria for identifying the items that anchor considers performance at more than one benchmark.

For multiple-choice items, a criterion of 65 percent was used for each benchmark being analyzed, since students would be likely (about two thirds of the time) to answer the item correctly. A criterion of less than 50 percent was used for the next lower benchmark, because with this response probability, students were more likely to have answered the item incorrectly than correctly. The criteria for each benchmark are outlined below.

- A multiple-choice item anchored at the Intermediate International Benchmark (475) if at least 65 percent of students scoring in the range answered the item correctly. Because this was the lowest benchmark described, there were no further criteria.
- A multiple-choice item anchored at the High International Benchmark (550) if at least 65 percent of students scoring in the range answered the item correctly, and less than 50 percent of students at the Intermediate International Benchmark answered the item correctly.
- A multiple-choice item anchored at the Advanced International Benchmark (625) if at least 65 percent of students scoring in the range answered the item correctly, and less than 50 percent of students at the High International Benchmark answered the item correctly.

To include all of the multiple-choice items in the anchoring process and provide information about content domains and cognitive processes that might not otherwise have had many anchor items, the concept of items that "almost anchored" was introduced. These were items that met slightly less stringent criteria for being answered correctly. The criteria to identify multiple-choice items that "almost anchored" were that at least 55 percent of students scoring in the range answered the item correctly and less than 50 percent of students at the next lowest benchmark answered correctly. To be completely inclusive for all items, items that met only the criterion that at least 55 percent of the students answered correctly (regardless of the performance of students at the next lower point) were also identified. The categories of items were mutually exclusive, and ensured that all of the items were available to inform the descriptions of student achievement at the anchor levels. A multiple-choice item was considered to be "too difficult" to anchor if less than 55 percent of students at the advanced benchmark answered the item correctly.

A somewhat less strict criterion was used for all the constructedresponse items, because students had much less scope for guessing. For constructed-response items, the criterion of 50 percent was used for the benchmark without any discrimination criterion for the next lower benchmark. A constructed-response item anchored at one of the international benchmarks if at least 50 percent of students at that benchmark answered the item correctly. A constructed-response item
was considered to be "too difficult" to anchor if less than 50 percent of students at the advanced benchmark answered the item correctly.

For students scoring in the range around each international benchmark, the percentage of those students that answered each item correctly was computed. To compute these percentages, students in each country were weighted to contribute proportional to the size of the student population in a country. Most of the TIMSS Advanced 2008 items were scored 1 point for a correct answer and o points for other answers. For these items, the percentage of students at each benchmark who answered each item correctly was computed. For the relatively few constructed-response items scored for partial or full credit, percentages were computed for the students receiving full credit. Then the criteria described above were applied to identify the items that anchored, almost anchored, and met only the 55 to 65 percent criteria. Exhibit 8.24 presents the number of advanced mathematics and physics items that anchored at each international benchmark.

Exhibit 8.24 Number of Items Anchoring at Each International Benchmark

| Intermediate (475) | High (550) | Advanced (625) | Too Difficult <br> to Anchor | Total |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Advanced Mathematics | 16 | 23 | 21 | 11 | 71 |
| Physics | 17 | 14 | 22 | 15 | 68 |

In preparation for the committee review, the advanced mathematics and physics items were organized into separate binders. The items were grouped by international benchmark and, within benchmark, the items were sorted by content area and then by the anchoring criteria they met: items that anchored, followed by items that almost anchored, followed by items that met only the 55 to 65 percent criteria. The following information was included for each item: content area, cognitive domain, maximum points, answer key, release
status, percent correct at each benchmark, and overall international percent correct. For constructed-response items, the scoring guides were included.

The TIMSS \& PIRLS International Study Center staff convened the committee for a three-day meeting in Boston to complete three tasks: The committee (1) worked through each item and arrived at a short description of the knowledge, understanding, or skills demonstrated by students who answered the item correctly; (2) developed a description (in detailed and summary form) of the level of advanced mathematics or physics proficiency demonstrated by students at each of the three international benchmarks to publish in the TIMSS Advanced 2008 International Report; and (3) selected example items that supported and illustrated the benchmark descriptions to publish together with the descriptions.

### 8.6 Capturing the Uncertainty in the TIMSS Advanced 2008 Student Achievement Scores

To obtain estimates of students' proficiency in advanced mathematics and physics that were both accurate and cost-effective, TIMSS Advanced 2008 made extensive use of probability sampling techniques to sample students from national student populations, and applied matrix sampling methods to target individual students with a subset of the entire set of assessment materials. Statistics computed from these student samples were used to estimate population parameters. This approach made efficient use of resources, in particular keeping student response burden to a minimum, but at a cost of some variance, or uncertainty, in the statistics. To quantify this uncertainty, each statistic in the TIMSS Advanced 2008 International Report (Mullis, Martin, Robitaille, \& Foy, 2009) is accompanied by an estimate of its standard error. These standard errors incorporate components reflecting the
uncertainty due to generalizing from student samples to the entire student populations (sampling variance), and to inferring students' performance on the entire assessment from their performance on the subset of items that they took (imputation variance).

### 8.6.1 Estimating the Sampling Variance

The TIMSS Advanced 2008 sample design applied a stratified multistage cluster-sampling technique to the problem of selecting efficient and accurate samples of students while working with schools and classes. This design capitalized on the structure of the student population (i.e., students grouped in classes within schools) to derive student samples that permitted efficient and economical data collection. Unfortunately, such a complex sample design complicates the task of computing standard errors to quantify sampling variability.

When, as in TIMSS Advanced, the sample design involves multi-stage cluster sampling, there are several options for estimating sampling errors that avoid the assumption of simple random sampling (Wolter, 1985). The jackknife repeated replication technique (JRR) was chosen by TIMSS because it is computationally straightforward and provides approximately unbiased estimates of the sampling errors of means, totals, and percentages.

The variation on the JRR technique used in TIMSS Advanced 2008 is described in Johnson and Rust (1992). It assumes that the primary sampling units (PSUs) can be paired in a manner consistent with the sample design, with each pair regarded as members of a pseudostratum for variance estimation purposes. When used in this way, the JRR technique appropriately accounts for the combined effect of the between- and within-PSU contributions to the sampling variance. The general use of JRR entails systematically assigning pairs of schools to sampling zones, and randomly selecting one of these schools to have
its contribution doubled and the other to have its contribution zeroed, so as to construct a number of "pseudo-replicates" of the original sample. The statistic of interest is computed once for the entire original sample, and once again for each jackknife pseudo-replicate sample. The variation between the estimates for each of the jackknife replicate samples and the original sample estimate is the jackknife estimate of the sampling error of the statistic.

### 8.6.2 Constructing Sampling Zones

To apply the JRR technique used in TIMSS Advanced 2008, successive sampled schools were paired and assigned to a series of groups known as sampling zones. This was done at Statistics Canada by working through the list of sampled schools in the order in which they were selected and assigning the first and second participating schools to the first sampling zone, the third and fourth participating schools to the second zone, and so on. A maximum of 75 zones were used, although most countries had fewer because they generally sampled less than 150 schools. When more than 75 zones were constructed, as was the case in Lebanon, they were collapsed to keep the total number to 75 .

Sampling zones were constructed within explicit strata. When there was an odd number of schools in an explicit stratum, either by design or because of school non-response, the students in the notpaired school were randomly divided to make up two "quasi" schools for the purposes of calculating the jackknife standard error. ${ }^{17}$ Each sampling zone then consisted of a pair of schools or "quasi" schools. Exhibit 8.25 shows the number of sampling zones in each country.

Within each sampling zone, each school was assigned an indicator $\left(u_{j}\right)$, coded randomly to o or 1 , such that one school had a value of zero, and the other a value of 1 . This indicator determined whether the weights for the sampled students in the school in this zone were to be doubled ( $u_{j}=1$ ) or zeroed ( $u_{j}=0$ ) for the purposes of creating the pseudo-replicate samples.

Exhibit 8.25 $\begin{aligned} & \text { Number of Sampling Zones Used in the TIMSS Advanced } 2008 \\ & \text { Countries }\end{aligned}$

| Country | Advanced Mathematics | Physics |
| :--- | :---: | :---: |
| Armenia | 55 | 55 |
| Italy | 46 | 46 |
| Iran, Islamic Rep. of | 60 | 60 |
| Lebanon | 75 | 75 |
| Netherlands | 56 | 58 |
| Norway | 55 | 52 |
| Philippines | 61 | - |
| Russian Federation | 45 | 47 |
| Slovenia | 53 | 66 |
| Sweden | 59 | 61 |

### 8.6.3 Computing the Sampling Variance Using the JRR Method

The formula for the sampling variance of a statistic $t$, based on the JRR algorithm used in TIMSS Advanced 2008, is given by the following equation:

$$
\operatorname{Var}_{j r r}(t)=\sum_{h=1}^{H}\left[t\left(J_{h}\right)-t(S)\right]^{2}
$$

where $H$ is the total number of sampling zones in the sample of the country under consideration. The term $t(S)$ corresponds to the statistic of interest for the whole sample computed with the overall sampling weights. ${ }^{18}$ The term $t\left(J_{h}\right)$ denotes the same statistic using the $h^{\text {th }}$ jackknife replicate sample $J_{h}$ and its set of replicate sampling weights, which are identical to the overall sampling weights, except for the students in the $h^{t h}$ sampling zone. In the $h^{\text {th }}$ zone, all students belonging to one of the randomly selected schools of the pair were removed, and the students belonging to the other school in the zone were included twice. In practice, this was accomplished by recoding to zero the sampling weights for the students in the school to be excluded
from the replication, and multiplying by 2 the sampling weights of the remaining students within the $h^{\text {th }}$ pair. Each sampled student was assigned a vector of 75 replicate sampling weights $W_{h i}$, where $h$ took values from 1 to 75 . If $W_{o i}$ was the overall sampling weight of student $i$, the $h$ replicate weights for that student were computed as:

$$
W_{h i}=W_{0 i} \cdot k_{h i}
$$

where

$$
k_{h i}=\left\{\begin{array}{cl}
2 \cdot u_{j} & \text { if student } i \text { is in school } j \text { of sampling zone } h \\
1 & \text { otherwise }
\end{array}\right\}
$$

The school-level indicators $u_{j}$ determined which students in a sampling zone would get zero weights and which ones would get double weights, on the basis of the school within the pair from which the students were sampled. The process of setting the $k_{h i}$ values for all sampled students and across all sampling zones is illustrated in Exhibit 8.26. Thus, the computation of the JRR variance estimate for any statistic in TIMSS Advanced 2008 required the computation of the statistic up to 76 times for any given country: once to obtain the statistic for the full sample based on the overall weights $W_{o i}$, and up to 75 times to obtain the statistics for each of the jackknife replicate samples $J_{h}$ using a set of replicate weights $W_{h i}$.

Exhibit 8.26 Construction of Replicate Weights Across Sampling Zones in TIMSS Advanced 2008

| Sampling Zone | $u_{j}$ | $k_{h i}$ for Computing JRR Replicate Sampling Weights |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | 1 | 2 | 3 | ... | h | $\ldots$ | 75 |
| 1 | 0 | 0 | 1 | 1 | $\ldots$ | 1 | $\ldots$ | 1 |
|  | 1 | 2 |  |  |  |  |  |  |
| 2 | 0 | 1 | 0 | 1 | $\ldots$ | 1 | $\ldots$ | 1 |
|  | 1 |  | 2 |  |  |  |  |  |
| 3 | 0 | 1 | 1 | 0 | $\ldots$ | 1 | ... | 1 |
|  | 1 |  |  | 2 |  |  |  |  |
| $\vdots$ | $\vdots$ | : | $\vdots$ | $\vdots$ | $\because$ | $\vdots$ | $\vdots$ | $\vdots$ |
| h | 0 | 1 | 1 | 1 | $\ldots$ | 0 | ... | 1 |
|  | 1 |  |  |  |  | 2 |  |  |
| $\vdots$ | : | $\vdots$ | $\vdots$ | $\vdots$ | $\ldots$ | $\vdots$ | $\because$ | : |
| 75 | 0 | 1 | 1 | 1 | ... | 1 | ... | 0 |
|  | 1 |  |  |  |  |  |  | 2 |

In the TIMSS Advanced 2008 analyses, 75 replicate weights were computed for each country regardless of the number of actual zones within the country. If a country had fewer than 75 zones, then the additional replicate weights where $h$ was greater than the number of zones within the country were all set equal to the overall sampling weight. Although this involved some redundant computations, having 75 replicate weights for each country had no effect on the magnitude of the error variance computed using the jackknife formula, and it simplified the computation of standard errors for numerous countries at the same time. All standard errors presented in the TIMSS Advanced 2008 international report were computed using SAS programs developed at the TIMSS \& PIRLS International Study Center.

### 8.6.4 Estimating the Imputation Variance

The TIMSS Advanced 2008 item pool was far too extensive to be administered in its entirety to any one student, and so a matrixsampling test design was developed whereby each student was given a single test booklet containing only a part of the entire assessment. ${ }^{19}$ The results for all of the booklets were then aggregated using item response theory to provide results for the entire assessment. Because each student responded to just a subset of the assessment items, it was necessary to use multiple imputation (the generation of plausible values) to derive reliable estimates of student performance on the assessment as a whole. Since every student's proficiency estimate incorporates some uncertainty arising from this imputation, TIMSS Advanced followed the customary procedure of generating five estimates for each student and using the variability among them as a measure of the imputation uncertainty, or error. In the TIMSS Advanced 2008 international report, the imputation error for each variable has been combined with the sampling error for that variable to provide a standard error that incorporates both.

The general procedure for estimating the imputation variance using plausible values is described in Mislevy, Beaton, Kaplan, and Sheenan (1992). First, compute the statistic $t$ for each set of $M$ plausible values. The statistics $t_{m}$, where $m=1,2, \ldots, M$, can be anything estimable from the data, such as a mean, the difference between means, percentiles, and so forth.

Once the statistics $t_{m}$ are computed, the imputation variance of the statistic $t$ is then calculated as:

$$
\operatorname{Var}_{i m p}(t)=\left(1+\frac{1}{M}\right) \operatorname{Var}\left(t_{1}, \ldots t_{M}\right)
$$

where $M$ is the number of plausible values used in the calculation, and $\operatorname{Var}\left(t_{p}, \ldots, t_{M}\right)$ is the usual variance of the $M$ estimates computed using each plausible value.

### 8.6.5 Combining the Sampling and Imputation Variance

The standard errors of all the proficiency statistics reported by TIMSS Advanced include both sampling and imputation variance components. These standard errors were computed using the following formula:

$$
\operatorname{Var}\left(t_{p v}\right)=\operatorname{Var}_{j r r}\left(t_{1}\right)+\operatorname{Var}_{i m p}(t)
$$

where $\operatorname{Var}_{j r r}\left(t_{1}\right)$ is the sampling variance computed for the first plausible value ${ }^{20}$ and $\operatorname{Var}_{i m p}(t)$ is the imputation variance. The TIMSS Advanced 2008 User Guide for the International Database (Foy \& Arora, 2009) contains programs in SAS and SPSS that compute each of these variance components for the TIMSS Advanced 2008 data. Furthermore, the IEA IDB Analyzer-software provided with the international database-automatically computes standard errors as described in this section.

Exhibit 8.27 shows basic summary statistics for overall advanced mathematics and physics achievement in the TIMSS Advanced 2008 assessments. It presents the student sample size, the mean and standard deviation averaged across the five plausible values, the jackknife sampling error for the mean, and the overall standard error for the mean, which includes the imputation error.

[^5]Exhibit 8.27 Summary Statistics and Standard Errors for Proficiency in TIMSS Advanced 2008

| Advanced Mathematics |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Country | Sample Size | Mean Proficiency | Standard Deviation | Jackknife Sampling Error | Overall Standard Error |
| Armenia | 858 | 432.760 | 95.466 | 3.090 | 3.675 |
| Iran, Islamic Rep. of | 2,425 | 496.750 | 98.767 | 6.306 | 6.369 |
| Italy | 2,143 | 448.779 | 95.468 | 6.870 | 7.142 |
| Lebanon | 1,612 | 544.726 | 60.472 | 2.258 | 2.318 |
| Netherlands | 1,537 | 552.470 | 45.797 | 1.917 | 2.647 |
| Norway | 1,932 | 439.224 | 86.910 | 4.679 | 4.990 |
| Philippines | 4,091 | 355.189 | 105.545 | 5.374 | 5.522 |
| Russian Federation | 3,185 | 560.984 | 90.972 | 7.138 | 7.213 |
| Slovenia | 2,156 | 457.316 | 84.850 | 3.890 | 4.151 |
| Sweden | 2,303 | 412.806 | 103.265 | 5.370 | 5.571 |


| Physics |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Country | Sample Size | Mean Proficiency | Standard Deviation | Jackknife Sampling Error | Overall Standard Error |
| Armenia | 894 | 495.067 | 100.284 | 5.125 | 5.363 |
| Iran, Islamic Rep. of | 2,434 | 459.856 | 115.728 | 7.145 | 7.204 |
| Italy | 1,861 | 422.238 | 102.558 | 7.510 | 7.624 |
| Lebanon | 1,595 | 443.542 | 78.324 | 2.615 | 2.990 |
| Netherlands | 1,511 | 582.474 | 53.723 | 3.547 | 3.703 |
| Norway | 1,640 | 534.142 | 78.147 | 4.182 | 4.212 |
| Russian Federation | 3,166 | 521.220 | 120.490 | 10.140 | 10.172 |
| Slovenia | 1,097 | 534.941 | 80.247 | 1.507 | 1.941 |
| Sweden | 2,291 | 496.950 | 91.865 | 5.509 | 5.651 |

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[^0]:    Keys: Diff: Percent correct score; Disc: Item discrimination; Pct_A...E: Percent choosing option; Pct_OM, NR: Percent Omitted and Not Reached; PB_A...E: Point Biserial for option; PB_OM: Point Biserial for Omittted. RDIFF= Rasch difficulty.

    Flags: A= Ability not ordered/Attractive distractor; $C=$ Difficulty less than chance; $D=$ Negative/low discrimination; E= Easier than average;

[^1]:    Keys: Diff: Percent correct score; Disc: Item discrimination; Pct_0...3: Percent obtaining score level; Pct_OM, NR: Percent Omitted and Not Reached; PB_0...3: Point Biserial for score level; PB_OM: Point Biserial for Omittted; RDIFF= Rasch difficulty ${ }^{\text {; }}$,
    Reliability: (Cases) Responses double scored; (Score) Percent agreement on score; (Code) Percent agreem
    $A=$ Ability not ordered/Attractive distractor; $C=$ Difficulty less than chance; $D=$ Negative/low discrimination; E= Easier than average;
    Flags:

[^2]:    V1 = Percent scoring 1 or better
    V2 = Percent scoring 2 or better
    Percent right for boys and girls corresponds to the percent obtaining the maximum score on the item. Because of missing gender information, some totals may appear inconsistent.

[^3]:    12 Because Denmark and Israel failed to satisfy the 1995 sampling guidelines, they were not included in the item calibrations for the rescaling, as was also the case for the original scaling.

[^4]:    14 Denmark and Israel, which had been excluded from the item calibrations, were included among the countries for which proficiency scores were produced.

[^5]:    20 Under ideal circumstances and with unlimited computing resources, the JRR sampling variance would be computed for each of the plausible values and the imputation variance as described here. This would require computing the same statistic up to 380 times (once overall for each of the five plausible values using the overall sampling weights, and then 75 times more for each plausible value using the complete set of replicate weights). An acceptable shortcut, however, is to compute the JRR sampling variance component using only one plausible value (the first one), and then the imputation variance using the five plausible values. Using this approach, a statistic needs to be computed only 80 times.

